

Distinguishing Distributions with Maximum Testing Power

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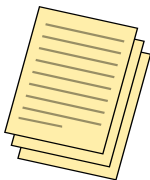
Realeyes, Budapest
August 24, 2016

- Motivating examples: NLP, computer vision.
- Two-sample test: t-test $\rightarrow$ distribution features.
- Linear-time, interpretable, high-power, nonparametric t-test.
- Numerical illustrations.

Motivating examples

Motivating example-1: NLP

- Given: two categories of documents (Bayesian inference, neuroscience).
- Task:
 - test their distinguishability,
 - most discriminative words $\rightarrow$ interpretability.



Motivating example-2: computer vision



- Given: two sets of faces (happy, angry).
- Task:
 - check if they are different,
 - determine the most discriminative features/regions.

Contribution:

- We propose a nonparametric t-test.
- It gives a reason why H_0 is rejected.
- It has high test power.
- It runs in linear time.

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- We propose a **nonparametric t-test**.
- It gives a **reason why H_0 is rejected**.
- It has **high test power**.
- It runs in **linear time**.

Dissemination, code:

- NIPS-2016 [Jitkrittum et al., 2016]: full oral = top 1.84%.
- <https://github.com/wittawatj/interpretable-test>.

Two-sample test, distribution features

What is a two-sample test?

- Given:

- $X = \{\mathbf{x}_i\}_{i=1}^n \stackrel{i.i.d.}{\sim} \mathbb{P}$, $Y = \{\mathbf{y}_j\}_{j=1}^n \stackrel{i.i.d.}{\sim} \mathbb{Q}$.
- Example: $\mathbf{x}_i = i^{th}$ happy face, $\mathbf{y}_j = j^{th}$ sad face.

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$$H_1 : \mathbb{P} \neq \mathbb{Q}.$$

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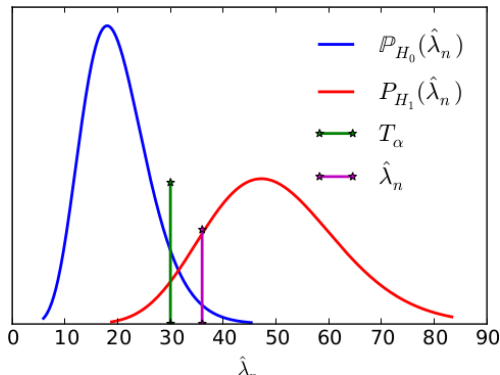
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- Assume $X, Y \subset \mathbb{R}^d$.

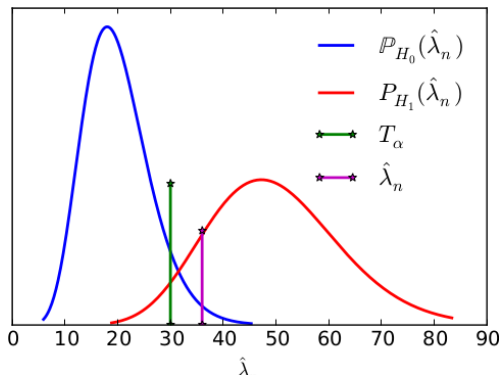
Ingredients of two-sample test

- Test statistic: $\hat{\lambda}_n = \hat{\lambda}_n(X, Y)$, random.
- Significance level: $\alpha = 0.01$.
- Under H_0 : $P_{H_0}(\underbrace{\hat{\lambda}_n \leq T_\alpha}_{\text{correctly accepting } H_0}) = 1 - \alpha$.



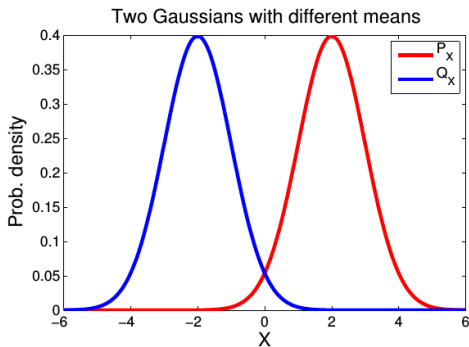
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- Under H_1 : $P_{H_1}(T_\alpha < \hat{\lambda}_n) = P(\text{correctly rejecting } H_0) =: \text{power}$.



Towards representations of distributions: EX

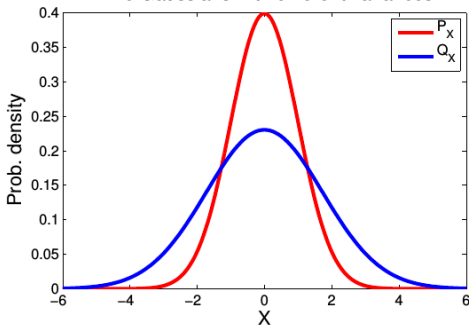
- Given: 2 Gaussians with different means.
- Solution: *t*-test.



Towards representations of distributions: $\mathbb{E}X^2$

- Setup: 2 Gaussians; same means, different variances.
- Idea: look at 2nd-order features of RVs.

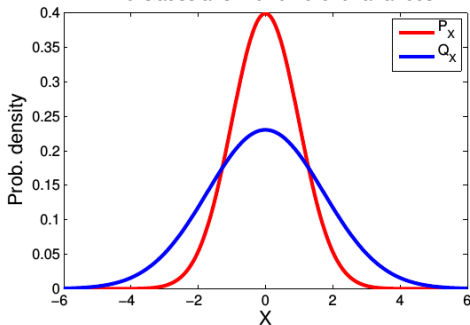
Two Gaussians with different variances



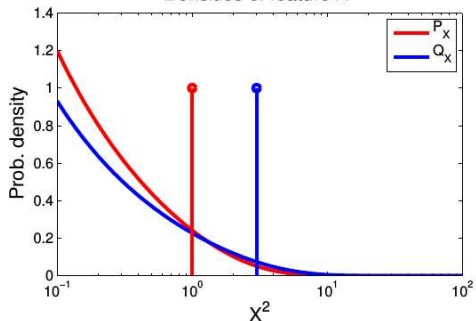
Towards representations of distributions: $\mathbb{E}X^2$

- Setup: 2 Gaussians; same means, different variances.
- Idea: look at 2nd-order features of RVs.
- $\varphi_X = x^2 \Rightarrow$ difference in $\mathbb{E}X^2$.

Two Gaussians with different variances

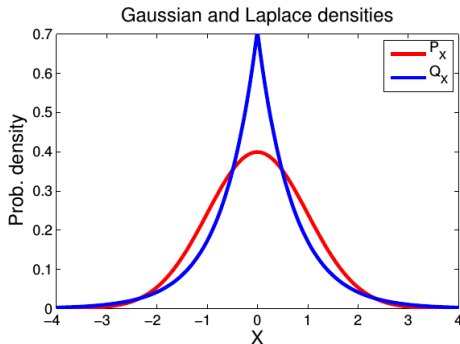


Densities of feature X^2



Towards representations of distributions: further moments

- Setup: a Gaussian and a Laplacian distribution.
- Challenge: their means *and* variances are the same.
- Idea: look at higher-order features.



Let us consider feature/distribution representations!

Kernel: similarity between features

- Given: $\mathbf{x}$ and $\mathbf{x}'$ objects (images or texts).

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Kernel: similarity between features

- Given: $\mathbf{x}$ and $\mathbf{x}'$ objects (images or texts).
- Question: how similar they are?
- Define **features** of the objects:

$\varphi_{\mathbf{x}}$: features of $\mathbf{x}$,

$\varphi_{\mathbf{x}'}$: features of $\mathbf{x}'$.

- **Kernel**: inner product of these features

$$k(\mathbf{x}, \mathbf{x}') := \langle \varphi_{\mathbf{x}}, \varphi_{\mathbf{x}'} \rangle .$$

Kernel examples on $\mathbb{R}^d$ ($\gamma > 0, p \in \mathbb{Z}^+$)

- Polynomial kernel:

$$k(\mathbf{x}, \mathbf{y}) = (\langle \mathbf{x}, \mathbf{y} \rangle + \gamma)^p.$$

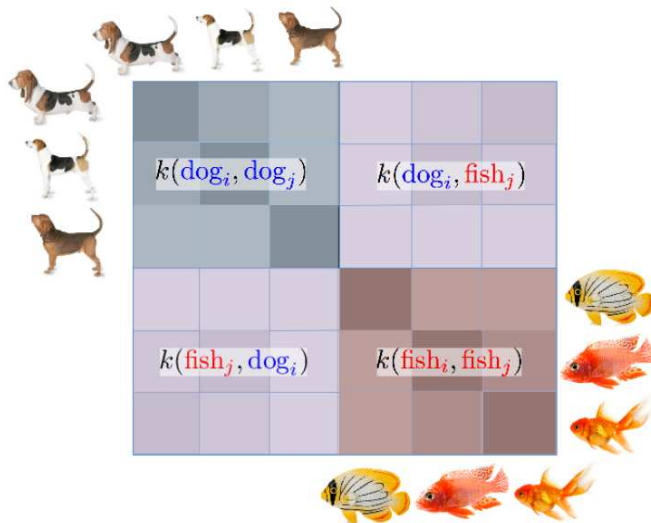
- Gaussian kernel:

$$k(\mathbf{x}, \mathbf{y}) = e^{-\gamma \|\mathbf{x} - \mathbf{y}\|_2^2}.$$

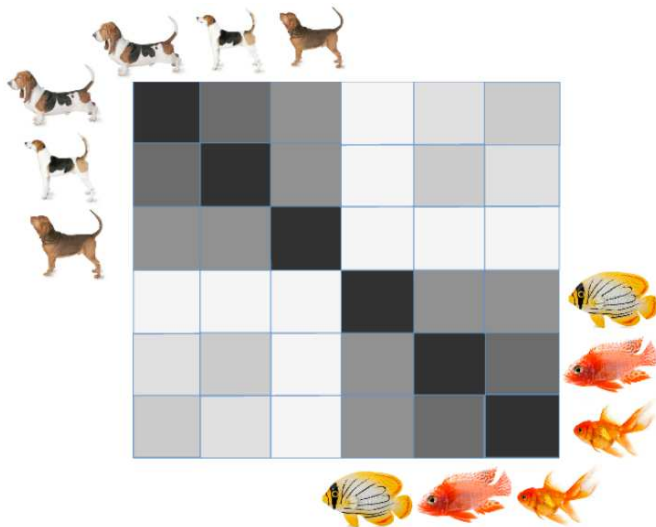
Towards distribution features



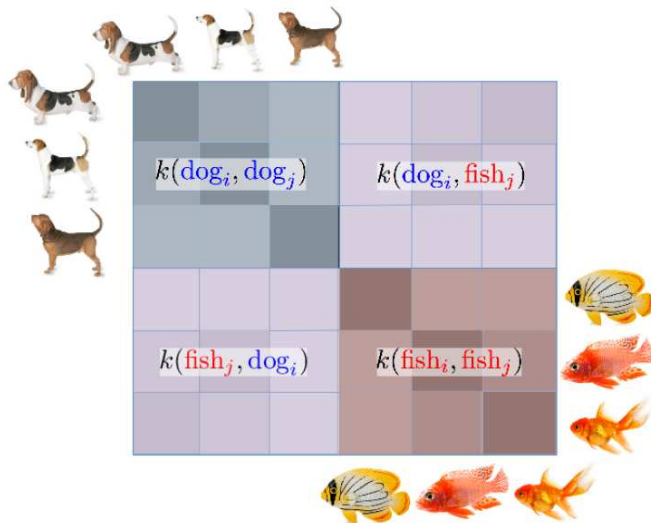
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Towards distribution features



Towards distribution features



$$\widehat{MMD}^2(\mathbb{P}, \mathbb{Q}) = \overline{K_{\mathbb{P}, \mathbb{P}}} + \overline{K_{\mathbb{Q}, \mathbb{Q}}} - 2\overline{K_{\mathbb{P}, \mathbb{Q}}} \quad (\text{without diagonals in } \overline{K_{\mathbb{P}, \mathbb{P}}}, \overline{K_{\mathbb{Q}, \mathbb{Q}}})$$

Kernel $\rightarrow$ distribution feature

- Kernel recall: $k(\mathbf{x}, \mathbf{x}') = \langle \varphi_{\mathbf{x}}, \varphi_{\mathbf{x}'} \rangle$.

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- Valid test [Gretton et al., 2012]. Challenges:
 - 1 Threshold choice: 'ugly' asymptotics of $n \widehat{MMD^2}(\mathbb{P}, \mathbb{P})$.
 - 2 Test statistic: quadratic time complexity.

Linear-time tests

Linear-time 2-sample test

- Recall:

$$MMD^2(\mathbb{P}, \mathbb{Q}) = \|\mu_{\mathbb{P}} - \mu_{\mathbb{Q}}\|_{\mathcal{H}(k)}^2.$$

- Changing [Chwialkowski et al., 2015] this to

$$\rho^2(\mathbb{P}, \mathbb{Q}) := \frac{1}{J} \sum_{j=1}^J [\mu_{\mathbb{P}}(\mathbf{v}_j) - \mu_{\mathbb{Q}}(\mathbf{v}_j)]^2.$$

with random $\{\mathbf{v}_j\}_{j=1}^J$ test locations

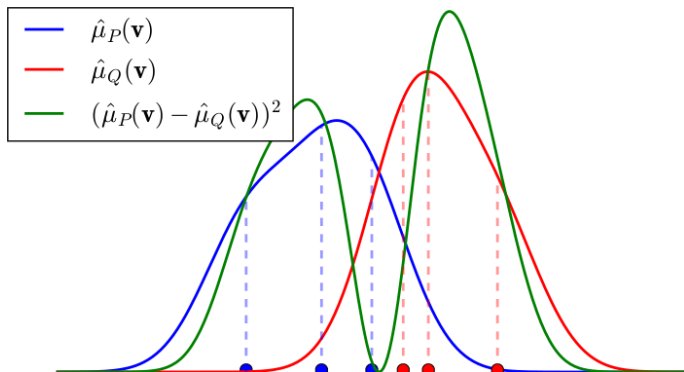
ρ is a metric (a.s.). How do we estimate it? Distribution under H_0 ?

Estimation

Estimate

$$\widehat{\rho^2(\mathbb{P}, \mathbb{Q})} = \frac{1}{J} \sum_{j=1}^J [\hat{\mu}_{\mathbb{P}}(\mathbf{v}_j) - \hat{\mu}_{\mathbb{Q}}(\mathbf{v}_j)]^2,$$

where $\hat{\mu}_{\mathbb{P}}(\mathbf{v}) = \frac{1}{n} \sum_{i=1}^n k(\mathbf{x}_i, \mathbf{v})$. Using $k(\mathbf{x}, \mathbf{v}) = e^{-\frac{\|\mathbf{x}-\mathbf{v}\|^2}{2\sigma^2}}$,



Estimation – continued

$$\begin{aligned}\widehat{\rho^2(\mathbb{P}, \mathbb{Q})} &= \frac{1}{J} \sum_{j=1}^J [\hat{\mu}_{\mathbb{P}}(\mathbf{v}_j) - \hat{\mu}_{\mathbb{Q}}(\mathbf{v}_j)]^2 \\ &= \frac{1}{J} \sum_{j=1}^J \left[\frac{1}{n} \sum_{i=1}^n k(\mathbf{x}_i, \mathbf{v}_j) - \frac{1}{n} \sum_{i=1}^n k(\mathbf{y}_i, \mathbf{v}_j) \right]^2 = \frac{1}{J} \sum_{j=1}^J (\bar{\mathbf{z}}_n)_j^2 = \frac{1}{J} \bar{\mathbf{z}}_n^T \bar{\mathbf{z}}_n,\end{aligned}$$

where $\bar{\mathbf{z}}_n = \frac{1}{n} \sum_{i=1}^n \underbrace{[k(\mathbf{x}_i, \mathbf{v}_j) - k(\mathbf{y}_i, \mathbf{v}_j)]}_{=: \mathbf{z}_i}^J_{j=1} \in \mathbb{R}^J$.

- Good news: estimation is linear in n !
- Bad news: intractable null distr. $= \sqrt{n} \widehat{\rho^2(\mathbb{P}, \mathbb{P})} \xrightarrow{w} \text{sum of } J \text{ correlated } \chi^2$.

Normalized version gives tractable null

- Modified test statistic:

$$\hat{\lambda}_n = n\bar{\mathbf{z}}_n^T \Sigma_n^{-1} \bar{\mathbf{z}}_n,$$

where $\Sigma_n = \text{cov}(\{\mathbf{z}_i\}_i)$.

- Under H_0 :
 - $\hat{\lambda}_n \xrightarrow{w} \chi^2(J)$. $\Rightarrow$ Easy to get the $(1 - \alpha)$ -quantile!

Our idea

- Until this point: test locations ($\mathcal{V}$) are fixed.
- Instead: choose $\theta = \{\mathcal{V}, \sigma\}$ to
maximize lower bound on the test power.

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Theorem (Lower bound on power)

For large n , test power $\geq L(\lambda_n)$; L : explicit function, increasing.

- Here,
 - $\lambda_n = n\boldsymbol{\mu}^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}$: population version of $\hat{\lambda}_n$.
 - $\boldsymbol{\mu} = \mathbb{E}_{\mathbf{xy}}[\mathbf{z}_1]$, $\boldsymbol{\Sigma} = \mathbb{E}_{\mathbf{xy}}[(\mathbf{z}_1 - \boldsymbol{\mu})(\mathbf{z}_1 - \boldsymbol{\mu})^T]$.

Convergence of the λ_n estimator

Training objective $\hat{\lambda}_n(X_{tr}, Y_{tr})$ converges to λ_n .

- But λ_n is unknown.
- Split (X, Y) into (X_{tr}, Y_{tr}) and (X_{te}, Y_{te}) . Use $\hat{\lambda}_n(X_{tr}, Y_{tr}) \approx \lambda_n$.

Theorem (Guarantee on objective approximation)

$$\left| \sup_{\mathcal{V}, \mathcal{K}} \bar{\mathbf{z}}_n^T (\boldsymbol{\Sigma}_n + \gamma_n)^{-1} \bar{\mathbf{z}}_n - \sup_{\mathcal{V}, \mathcal{K}} \boldsymbol{\mu}^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu} \right| = \mathcal{O}\left(n^{-\frac{1}{4}}\right).$$

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Examples:

$$\mathcal{K} = \{k_\sigma(\mathbf{x}, \mathbf{y}) = e^{-\|\mathbf{x}-\mathbf{y}\|^2} : \sigma > 0\},$$

$$\mathcal{K} = \{k_{\mathbf{A}}(\mathbf{x}, \mathbf{y}) = e^{-(\mathbf{x}-\mathbf{y})^T \mathbf{A} (\mathbf{x}-\mathbf{y})} : \mathbf{A} \succ 0\}.$$

Numerical demos

Parameter settings

- Gaussian kernel (σ). $\alpha = 0.01$. $J = 1$. Repeat 500 trials.
- Report

$$\mathbb{P}(\text{reject } H_0) \approx \frac{\text{\#times } \hat{\lambda}_n > T_\alpha \text{ holds}}{\text{\#trials}}.$$

- Compare 4 methods
 - **ME-full**: Optimize $\mathcal{V}$ and Gaussian bandwidth σ .
 - **ME-grid**: Optimize σ . Fix $\mathcal{V}$ [Chwialkowski et al., 2015].
 - **MMD-quad**: Test with quadratic-time MMD [Gretton et al., 2012].
 - **MMD-lin**: Test with linear-time MMD [Gretton et al., 2012].
- Optimize kernels to power in MMD-lin, MMD-quad.

NLP: discrimination of document categories

- 5903 NIPS papers (1988-2015).
- Keyword-based category assignment into 4 groups:
 - Bayesian inference, Deep learning, Learning theory, Neuroscience
- $d = 2000$ nouns. TF-IDF representation.

Problem	n^{te}	ME-full	ME-grid	MMD-quad	MMD-lin
1. Bayes-Bayes	215	.012	.018	.022	.008
2. Bayes-Deep	216	.954	.034	.906	.262
3. Bayes-Learn	138	.990	.774	1.00	.238
4. Bayes-Neuro	394	1.00	.300	.952	.972
5. Learn-Deep	149	.956	.052	.876	.500
6. Learn-Neuro	146	.960	.572	1.00	.538

- Performance of ME-full [$\mathcal{O}(n)$] is comparable to MMD-quad [$\mathcal{O}(n^2)$].

NLP: most/least discriminative words

- Aggregating over trials; example: 'Bayes-Neuro'.
- Most discriminative words:
 - spike, markov, cortex, dropout, recurr, iii, gibb.
 - learned test locations: highly interpretable,
 - 'markov', 'gibb' ($\Leftarrow$ Gibbs): Bayesian inference,
 - 'spike', 'cortex': key terms in neuroscience.

NLP: most/least discriminative words

- Aggregating over trials; example: 'Bayes-Neuro'.
- Least discriminative ones:
circumfer, bra, dominiqu, rhino, mitra, kid, impostor.

Distinguish positive/negative emotions

- Karolinska Directed Emotional Faces (KDEF) [Lundqvist et al., 1998].
- 70 actors = 35 females and 35 males.
- $d = 48 \times 34 = 1632$. Grayscale. Pixel features.



Problem	n^{te}	ME-full	ME-grid	MMD-quad	MMD-lin
$\pm$ vs. $\pm$	201	.010	.012	.018	.008
$+$ vs. $-$	201	.998	.656	1.00	.578



- Learned test location (averaged) =

- We proposed a nonparametric t-test:
 - linear time,
 - high-power ($\approx$ 'MMD-quad'),
- 2 demos: discriminating
 - documents of different categories,
 - positive/negative emotions.

Thank you for the attention!



Acknowledgements: This work was supported by the Gatsby Charitable Foundation.

- Non-convexity, informative features.
- Number of locations (J).
- Computational complexity.
- Estimation of MMD^2 .

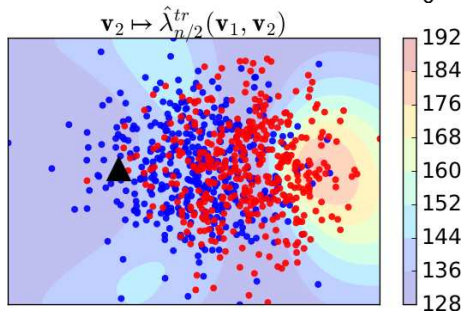
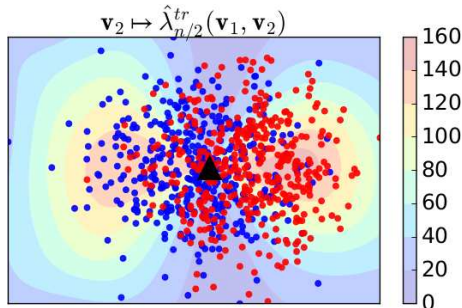
Non-convexity, informative features

- 2D problem:

$$\mathbb{P} := \mathcal{N}([0; 0], \mathbf{I}),$$

$$\mathbb{Q} := \mathcal{N}([1; 0], \mathbf{I}).$$

- $\mathcal{V} = \{\mathbf{v}_1, \mathbf{v}_2\}$.
- Fix $\mathbf{v}_1$ to $\blacktriangle$.
- Contour plot of $\mathbf{v}_2 \mapsto \hat{\lambda}_n(\{\mathbf{v}_1, \mathbf{v}_2\})$.



Number of locations (J)

- Small J :
 - often enough to detect the difference of $\mathbb{P}$ & $\mathbb{Q}$.
 - few distinguishing regions to reject H_0 .
 - faster test.
- Very large J :
 - test power need not increase monotonically in J (more locations $\Rightarrow$ statistic can gain in variance).
 - defeats the purpose of a linear-time test.

- Optimization & testing: linear in n .
- Testing: $\mathcal{O}(ndJ + nJ^2 + J^3)$.
- Optimization: $\mathcal{O}(ndJ^2 + J^3)$ per gradient ascent.

Squared difference between feature means:

$$\begin{aligned}MMD^2(\mathbb{P}, \mathbb{Q}) &= \|\mu_{\mathbb{P}} - \mu_{\mathbb{Q}}\|_{\mathcal{H}}^2 = \langle \mu_{\mathbb{P}} - \mu_{\mathbb{Q}}, \mu_{\mathbb{P}} - \mu_{\mathbb{Q}} \rangle_{\mathcal{H}} \\&= \langle \mu_{\mathbb{P}}, \mu_{\mathbb{P}} \rangle_{\mathcal{H}} + \langle \mu_{\mathbb{Q}}, \mu_{\mathbb{Q}} \rangle_{\mathcal{H}} - 2 \langle \mu_{\mathbb{P}}, \mu_{\mathbb{Q}} \rangle_{\mathcal{H}} \\&= \mathbb{E}_{\mathbb{P}, \mathbb{P}} k(\mathbf{x}, \mathbf{x}') + \mathbb{E}_{\mathbb{Q}, \mathbb{Q}} k(\mathbf{y}, \mathbf{y}') - 2 \mathbb{E}_{\mathbb{P}, \mathbb{Q}} k(\mathbf{x}, \mathbf{y}).\end{aligned}$$

Estimation of MMD^2

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Unbiased empirical estimate for $\{\mathbf{x}_i\}_{i=1}^n \sim \mathbb{P}$, $\{\mathbf{y}_j\}_{j=1}^n \sim \mathbb{Q}$:

$$\widehat{MMD^2}(\mathbb{P}, \mathbb{Q}) = \overline{K_{\mathbb{P}, \mathbb{P}}} + \overline{K_{\mathbb{Q}, \mathbb{Q}}} - 2\overline{K_{\mathbb{P}, \mathbb{Q}}}.$$

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