

Towards Outlier-Robust Statistical Inference on Kernel-Enriched Domains

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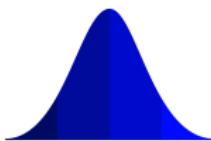
- RKHS: $\mathcal{H}_K = \overline{\{\sum_{i=1}^n \alpha_i K(\cdot, x_i)\}} \subset \mathbb{R}^{\mathcal{X}}$. $\varphi(x) = \underbrace{K(\cdot, x)}_{\text{A blue bell-shaped curve}} \in \mathcal{H}_K.$



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- We represent distributions in RKHSs: $\mu_{\mathbb{P}} := \int_{\mathcal{X}} \varphi(x) d\mathbb{P}(x) \in \mathcal{H}_K$.

- **Trees** [Collins and Duffy, 2001, Kashima and Koyanagi, 2002], **time series** [Cuturi, 2011], **strings** [Lodhi et al., 2002],
- **mixture models**, **hidden Markov models** or **linear dynamical systems** [Jebara et al., 2004],
- **sets** [Haussler, 1999, Gärtner et al., 2002], **fuzzy domains** [Guevara et al., 2017], **distributions** [Hein and Bousquet, 2005, Martins et al., 2009, Muandet et al., 2011],
- **groups** [Cuturi et al., 2005] $\xrightarrow{\text{spec.}}$ **permutations** [Jiao and Vert, 2018],
- **graphs** [Vishwanathan et al., 2010, Kondor and Pan, 2016].

Back to mean embeddings: $\mu_{\mathbb{P}}$

Very natural representation

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Trick

φ : on any kernel-endowed domain!

Mean embedding (\exists)

- $\mu_{\mathbb{P}} = \int_{\mathcal{X}} \varphi(x) d\mathbb{P}(x)$ exists $\Leftrightarrow \int_{\mathcal{X}} \underbrace{\|\varphi(x)\|_{\mathcal{H}_K}}_{\sqrt{K(x,x)}} d\mathbb{P}(x) < \infty.$

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Until now

We have defined $\mu_{\mathbb{P}}$ and $\text{MMD}(\mathbb{P}, \mathbb{Q})$.

- Applications:

- two-sample testing [Borgwardt et al., 2006, Gretton et al., 2012],
 - domain adaptation [Zhang et al., 2013], -generalization [Blanchard et al., 2017],
 - kernel Bayesian inference [Song et al., 2011, Fukumizu et al., 2013]
 - approximate Bayesian computation [Park et al., 2016], probabilistic programming [Schölkopf et al., 2015],
 - model criticism [Lloyd et al., 2014, Kim et al., 2016], goodness-of-fit [Balasubramanian et al., 2017],
 - distribution classification [Muandet et al., 2011, Lopez-Paz et al., 2015], [Zaheer et al., 2017], distribution regression [Szabó et al., 2016], [Law et al., 2018],
 - topological data analysis [Kusano et al., 2016].
- Review [Muandet et al., 2017].

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$$\stackrel{\text{or}}{=} \frac{1}{N(N-1)} \sum_{\substack{i,j \in [N] \\ i \neq j}} [K(x_i, x_j) + K(y_i, y_j)] - \frac{2}{N^2} \sum_{i,j \in [N]} K(x_i, y_j).$$

Goal of our work

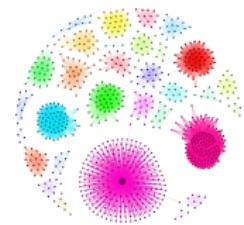
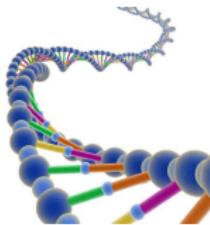
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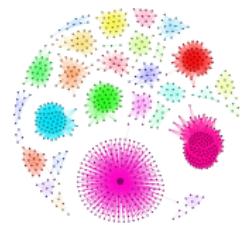
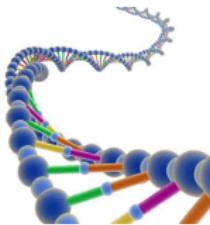


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Issue with average

A single outlier can ruin it.

Existing work

- Robust KDE [Kim and Scott, 2012]:

$$\mu_{\mathbb{P}} = \arg \min_{f \in \mathcal{H}_K} \int_{\mathcal{X}} \|f - K(\cdot, x)\|_{\mathcal{H}_K}^2 d\mathbb{P}(x)$$

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- **Consistency** : For finiteD features [Sinova et al., 2018]

$$\hat{\mu}_{\mathbb{P}, N, L} \xrightarrow{N \rightarrow \infty} \mu_{\mathbb{P}, L}. \quad (\text{empirical M-estimator in } \mathbb{R}^d)$$

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- ② Compute average in each block:

$$a_1 = \frac{1}{|S_1|} \sum_{i \in S_1} x_i, \dots, a_Q = \frac{1}{|S_Q|} \sum_{i \in S_Q} x_i.$$

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- ③ Estimate $\mathbb{E}X$: $\text{med}_{q \in [Q]} a_q$.

Idea on MMD (mean embedding: similarly)

- ① Use the IPM representation:

$$\text{MMD}(\mathbb{P}, \mathbb{Q}) = \sup_{f \in \mathcal{B}_K} \langle f, \mu_{\mathbb{P}} - \mu_{\mathbb{Q}} \rangle_{\mathcal{H}_K}.$$

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What can we show about this MONK estimator?

Assumptions:

- ① The # of samples contaminated can be (almost) half of the # of blocks :

$$\{(x_{n_j}, y_{n_j})\}_{j=1}^{N_c}, \quad N_c \leq Q(1/2 - \delta), \quad \delta \in (0, 1/2].$$

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Minimal 2nd-order condition .

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Minimal 2nd-order condition. Note: $\|A\| \leq \|A\|_{HS} \stackrel{(*)}{\leq} \|A\|_1$.

Consistency & outlier-robustness of $\widehat{\text{MMD}}_Q(\mathbb{P}, \mathbb{Q})$

Then, for any $\eta \in (0, 1)$ such that $Q = 72\delta^{-2} \ln(1/\eta)$ satisfies $Q \in (N_c / (\frac{1}{2} - \delta), N/2)$, with probability at least $1 - \eta$

$$\begin{aligned} & \left| \widehat{\text{MMD}}_Q(\mathbb{P}, \mathbb{Q}) - \text{MMD}(\mathbb{P}, \mathbb{Q}) \right| \\ & \leq \frac{12 \max \left(\sqrt{\frac{(\|\Sigma_{\mathbb{P}}\| + \|\Sigma_{\mathbb{Q}}\|) \ln(1/\eta)}{\delta N}}, 2\sqrt{\frac{\text{Tr}(\Sigma_{\mathbb{P}}) + \text{Tr}(\Sigma_{\mathbb{Q}})}{N}} \right)}{\delta}. \end{aligned}$$

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Discussion:

- ① N -dependence: $\mathcal{O}\left(\frac{1}{\sqrt{N}}\right)$ is optimal for MMD estimation [Tolstikhin et al., 2016].

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- Optimal sub-Gaussian deviation bound for **mean** estimation under minimal 2nd-order condition even on \mathbb{R}^d
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- They rely on tournament procedure: numerically hard.

Consistency & outlier-robustness of $\widehat{\text{MMD}}_Q(\mathbb{P}, \mathbb{Q})$

Then, for any $\eta \in (0, 1)$ such that $Q = 72\delta^{-2} \ln(1/\eta)$ satisfies $Q \in (N_c / (\frac{1}{2} - \delta), N/2)$, with probability at least $1 - \eta$

$$\left| \widehat{\text{MMD}}_Q(\mathbb{P}, \mathbb{Q}) - \text{MMD}(\mathbb{P}, \mathbb{Q}) \right| \leq \frac{12 \max \left(\sqrt{\frac{(\|\Sigma_{\mathbb{P}}\| + \|\Sigma_{\mathbb{Q}}\|) \ln(1/\eta)}{\delta N}}, 2\sqrt{\frac{\text{Tr}(\Sigma_{\mathbb{P}}) + \text{Tr}(\Sigma_{\mathbb{Q}})}{N}} \right)}{\delta}.$$

Discussion:

② Σ -dependence:

- Optimal sub-Gaussian deviation bound for **mean** estimation under minimal 2nd-order condition even on \mathbb{R}^d [Lugosi and Mendelson, 2019] – long-lasting open question.
- They rely on tournament procedure: numerically hard.
- Most **practical** convex relaxation [Hopkins, 2018]: $\mathcal{O}(N^{24})$.

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Discussion:

③ δ -dependence:

- Larger δ means less outliers,
 - the bound becomes tighter,
 - one needs less blocks.

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Discussion:

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- optimal?

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Discussion:

④ **breakdown point** – asymptotic concept:

- median \Rightarrow Using Q blocks is resistant to $Q/2$ outliers.
- Q can grow with N , as (almost) $N/2$.
- Breakdown point can be 25%.

Then, for any $\eta \in (0, 1)$ such that $Q = 72\delta^{-2} \ln(1/\eta)$ satisfies $Q \in (N_c / (\frac{1}{2} - \delta), N/2)$, with probability at least $1 - \eta$

$$\begin{aligned} & \left| \widehat{\text{MMD}}_Q(\mathbb{P}, \mathbb{Q}) - \text{MMD}(\mathbb{P}, \mathbb{Q}) \right| \\ & \leqslant \frac{12 \max \left(\sqrt{\frac{(\|\Sigma_{\mathbb{P}}\| + \|\Sigma_{\mathbb{Q}}\|) \ln(1/\eta)}{\delta N}}, 2\sqrt{\frac{\text{Tr}(\Sigma_{\mathbb{P}}) + \text{Tr}(\Sigma_{\mathbb{Q}})}{N}} \right)}{\delta}. \end{aligned}$$

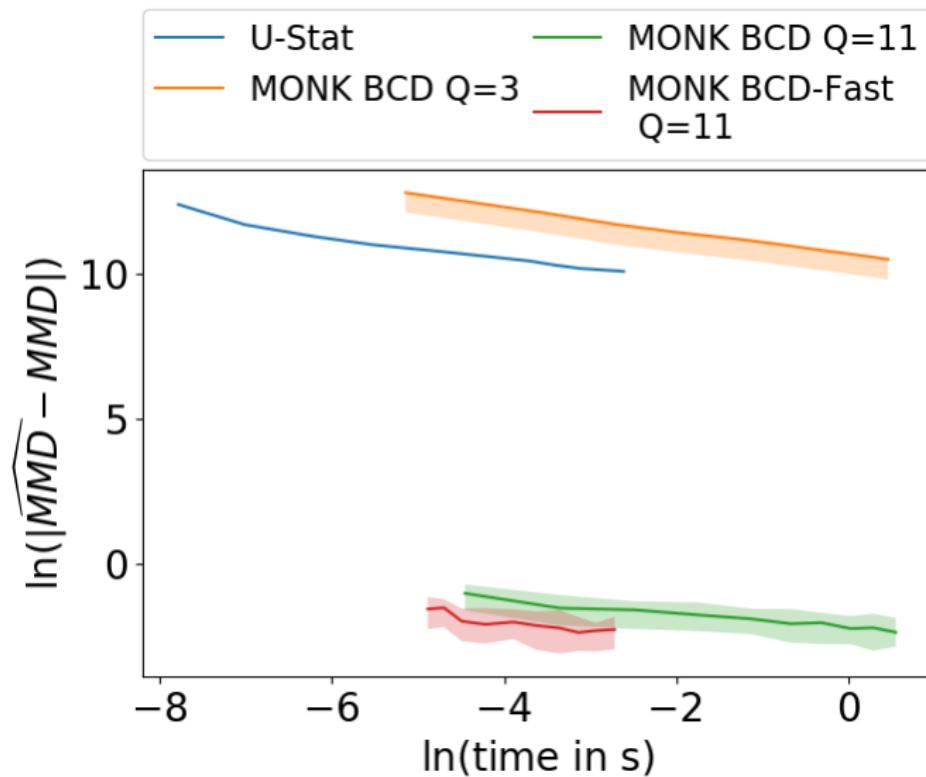
Discussion:

⑤ Unknown Q :

- One choose Q adaptively by the Lepski method.
- Same guarantee but with increased computational cost.

- ① No outliers / bounded kernel: MONK is a safe alternative.
- ② Relevant case: outliers & unbounded kernel.
 - $\mathbb{P} := \mathcal{N}(\mu_1, \sigma_1^2) \neq \mathbb{Q} := \mathcal{N}(\mu_2, \sigma_2^2)$. $\mu_m, \sigma_m \sim U[0, 1]$, fixed.
 - $N \in \{200, 400, \dots, 2000\}$.
 - 5-5 corrupted samples: $(x)_n^N_{n=N-4} = 2000$, $(y_n)_n^N_{n=N-4} = 4000$.
 - $(\mathbb{P}, \mathbb{Q}, K)$: MMD(\mathbb{P}, \mathbb{Q}) is analytic.
 - Performance:
 - 100 MC simulations,
 - median and quartiles.

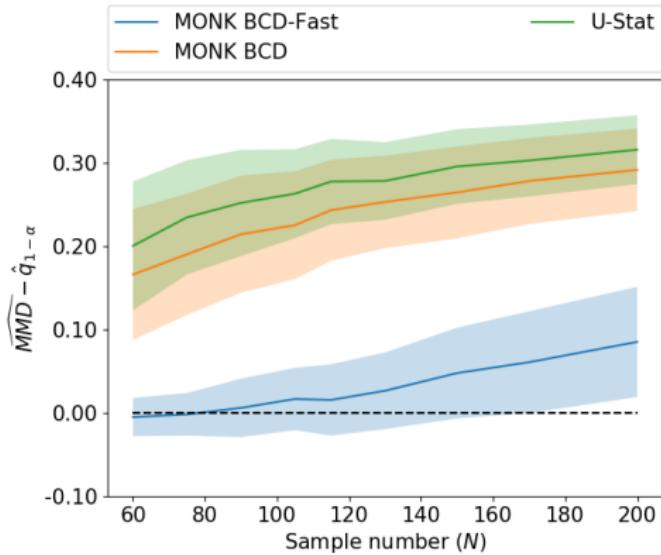
Numerical demo: quadratic kernel, $N_c = 5$ outliers



- Discrimination of 2 DNA categories (EI, IE).
- Subsequent String Kernel (K).
- Significance level: $\alpha = 0.05$.
- Performance:
 - 4000 MC simulations,
 - mean \pm std of $\widehat{\text{MMD}} - \hat{q}_{1-\alpha}$.
- $\hat{q}_{1-\alpha}$: Using 150 bootstrap permutations.

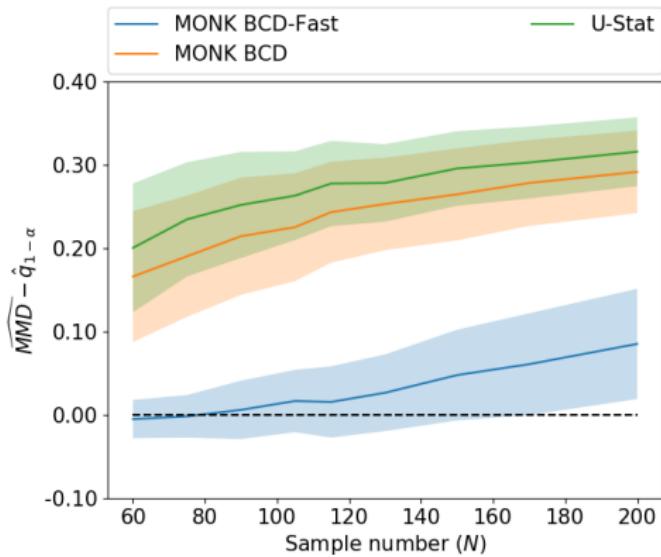
DNA analysis: plots

Inter-class: EI-IE

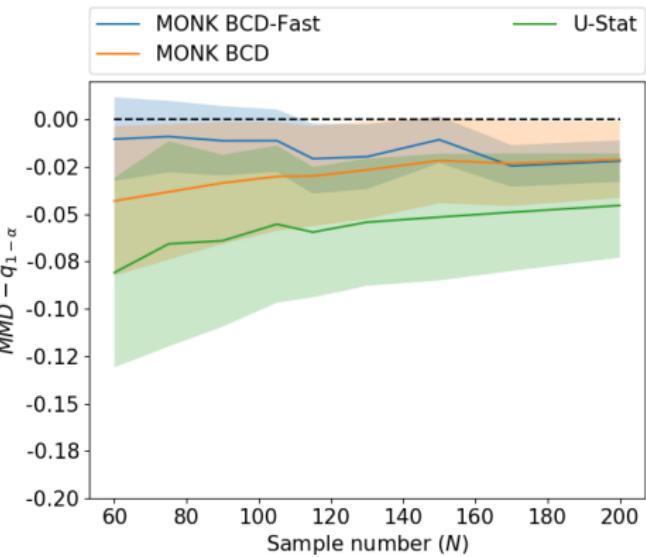


DNA analysis: plots

Inter-class: EI-IE,



Intra-class: EI-EI (IE-IE)



Summary

- Focus: Outlier-robust mean embedding & MMD estimation.
- Technique: median-of-means.
- Finite-sample guarantees (optimality), excessive resistance to contamination.

Summary

- Focus: Outlier-robust mean embedding & MMD estimation.
- Technique: median-of-means.
- Finite-sample guarantees (optimality), excessive resistance to contamination.
- Preprint, code:

MONK – Outlier-Robust Mean Embedding Estimation by
Median-of-Means, TR

(<http://arxiv.org/abs/1802.04784>).

<https://bitbucket.org/TimotheeMathieu/monk-mmd>

Thank you for the attention!



Computational complexity of MMD estimators

N : sample number, Q : number of blocks, T : number of iterations.

Method	Complexity
U-Stat	$\mathcal{O}(N^2)$
MONK BCD	$\mathcal{O}(N^3 + T [N^2 + Q \log(Q)])$
MONK BCD-Fast	$\mathcal{O}\left(\frac{N^3}{Q^2} + T \left[\frac{N^2}{Q} + Q \log(Q)\right]\right)$

Pseudo-code: 2-sample testing

Input: Two samples: $(X_n)_{n \in [N]}, (Y_n)_{n \in [N]}$. Number of bootstrap permutations: $B \in \mathbb{Z}^+$. Level of the test: $\alpha \in (0, 1)$. Kernel function with hyperparameter $\theta \in \Theta$: K_θ .

Split the dataset randomly into 3 equal parts:

$$[N] = \bigcup_{i=1}^3 I_i, \quad |I_1| = |I_2| = |I_3|.$$

Tune the hyperparameters using the 1st part of the dataset:

$$\hat{\theta} = \arg \max_{\theta \in \Theta} J_\theta := \widehat{\text{MMD}}_\theta((X_n)_{n \in I_1}, (Y_n)_{n \in I_1}).$$

Estimate the $(1 - \alpha)$ -quantile of $\widehat{\text{MMD}}_{\hat{\theta}}$ under the null, using B bootstrap permutations from $(X_n)_{n \in I_2} \cup (Y_n)_{n \in I_2}$: $\hat{q}_{1-\alpha}$.

Compute the test statistic on the third part of the dataset:

$$T_{\hat{\theta}} = \widehat{\text{MMD}}_{\hat{\theta}}((X_n)_{n \in I_3}, (Y_n)_{n \in I_3}).$$

Output: $T_{\hat{\theta}} - \hat{q}_{1-\alpha}$.

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